

# Social media, sentiment and public opinions: Evidence from #Brexit and #USElection

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**Abstract:** This paper studies information diffusion in social media and the role of bots in shaping public opinions. Using Twitter data on the 2016 E.U. Referendum (“Brexit”) and the 2016 U.S. Presidential Election, we find that diffusion of information on Twitter is largely complete within 1-2 hours. Stronger interactions across agents with similar beliefs are consistent with the “echo chambers” view of social media. Bots have a tangible effect on the tweeting activity of humans but the degree of bots’ influence depends on whether bots provide information consistent with humans’ priors. Overall, our results suggest that the aggressive use of Twitter bots, coupled with the fragmentation of social media and the role of sentiment, could contribute to the vote outcomes.

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## 1. Introduction

The rise of Internet has changed the way people communicate and acquire information. Recent years have witnessed a decline in the traditional news media consumption (Stempel et al., 2000) while the number of online news users soars (Gottfried and Shearer, 2016; Bialik and Matsa, 2017). Among different types of Internet-based media, social network has become an increasingly important information source for many people. Through social media, individuals can have instant and open access to news and narratives and can build networks to interact and share opinions. Key questions are how this communication revolution has influenced information flows across individuals and how one can influence these flows.

We attempt to answer this question by examining information dissemination in social media and using recent developments in the U.K. (2016 E.U. Referendum) and the U.S. (2016 Presidential Election) as two natural experiments. These two events were extremely high-profile so that people paid as much attention to these as one may reasonably expect. In addition, these experiments are “clean” in the sense that platforms in these plebiscites were diametrically opposed, which allows us to separate providers of information as well as consumers of information.

In this exercise, we identify sources of information that might have been used to shape public opinions. Specifically, we study two types of social media agents: real (“human”) users and social bots, computer algorithms used to produce automated content. Bots can create new bits of information and a sense of consensus in the society that is favorable for a given candidate/outcome. We use these bots as a source of variation in people’s information sets and see how this information influences “humans”, how it is spread across “humans”, and how the sentiment (tonality) of bots’ messages affects “humans”.

Data employed in this study were collected using Twitter Streaming Application Programming Interface (API).<sup>1</sup> Twitter is one of the most popular microblogging platforms with a significant number of users. For example, as of 2016, the number of U.K. Twitter users is estimated at 15.8 million while the number of U.S. Twitter users is about 67 million (Benoit, 2017). Given this popularity, Twitter generates an enormous quantity of legally available data for research.<sup>2</sup>

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<sup>1</sup> Twitter streaming API is a developer tool that allows collecting a random sample of real-time tweets with pre-defined attributes (e.g. keywords, usernames, or hashtags). See for details [https://developer.twitter.com/en/docs/tweets/filter-realtime/guides/powertrack\\_rules\\_and\\_filtering](https://developer.twitter.com/en/docs/tweets/filter-realtime/guides/powertrack_rules_and_filtering), accessed on April 27, 2018.

<sup>2</sup> In contrast to Cambridge Analytica/Facebook case, our data are collected directly from Twitter and the collection process does not breach any terms and condition of Twitter development tools.

These data include records of the users, tweets, and their metadata that allow us to track tweeting activities of different types of Twitter agents (bots and humans). Additionally, compared to other social network sites, Twitter's connections are more about connecting with people for information sharing purposes rather than personal interactions (Gruzd et al., 2011). Hence, one would expect to see clear information flows as well as different information clusters during important political events on Twitter. Furthermore, Twitter is used by people with extreme views whose voices are often invisible in the mainstream media. Thus, data from Twitter could give us a broader picture about public opinions during the EU Referendum in the U.K. and the 2016 U.S. Presidential Election beyond what we can observe from the traditional media.

We find that information about the Brexit and the 2016 U.S. Presidential Election is disseminated and absorbed among Twitter users within 50-70 minutes. This suggests that information rigidity could be very low for critically important issues with wide coverage or that news cycles in social media are short-lived. We also observe the differential impact of tweeting activities by user type. For example, "remain" supporters in the Brexit Referendum respond stronger and faster to messages created by other "remain" supporters when compared with the reaction to messages from "leave" supporters. Furthermore, human tweeting activity could be influenced by bots. The degree of influence depends on whether a bot provides information consistent with the priors of a human. For instance, a bot supporting the "leave" campaign has a stronger impact on a "leave" supporter than a "remain" supporter. Similarly, Trump supporters are more likely to react to messages spread by pro-Trump bots. Further examination shows that the sentiment of tweets plays an important role in how information is spread: a message with positive (negative) sentiment generates another message with the same sentiment. These results provide evidence consistent with the "echo chambers" effect in social media, that is, people tend to select themselves into groups of like-minded people so that their beliefs are reinforced while information from outsiders might be ignored. Therefore, social media platforms like Twitter could enhance ideological segmentation and make information more fragmented rather than more uniform across people. Finally, we provide a quantitative assessment of how bots' traffic contributed to the actual vote outcomes. Our results suggest that, given narrow margins of victories in each vote, bots' effect was likely marginal but possibly large enough to affect the outcomes.

This study is related to several strands of research. The first strand has assessed the influence of news media on the economic and political outcomes. For example, Tetlock (2007), Engelberg and

Parsons (2011), and Chen et al. (2014) show that media coverage of a company's stock is significantly related to future stock prices and trading volume. Media exposure could also have impacts on political outcomes or voter behaviors. Gerber et al. (2009) find that subscription to either Washington Post or Washington Times increases the support for Democratic candidates. Similarly, DellaVigna and Kaplan (2007) suggest that the introduction of Fox News channel increases the vote shares of Republican candidates as well as increases voter turnout. In contrast, the introduction of television which has less political coverage leads to a decline in political knowledge and thus turnout rates (Gentzkow, 2006). We contribute to this literature by examining information flows in online media. To the best of our knowledge, our paper is among very few studies that empirically investigate the influence of information consumption in social media on political outcomes.

The strong impact of news content documented in the first strand has motivated the second strand which examines the motivations of news and information consumption. Some papers suggest the ideological segregation that people tend to select themselves into group of like-minded people so their beliefs are reinforced (see, e.g., Mullainathan and Shleifer, 2005; Halberstam and Knight, 2016). This view holds for political communication when individuals are more likely to expose to news and information consistent with their political views (Adamic and Glance, 2005; Garrett, 2009a; Gruzd and Roy, 2014; Hong and Kim, 2016). However, Gentzkow and Shapiro (2011) suggest that although ideological segregation does exist in both online and offline news consumption, the level of polarization is low. In other words, exposure to information reinforcing individuals' pre-existing views does not shield them from following information that they disagree with (Hargittai et al., 2008; Garrett, 2009b; DellaVigna et al., 2014). Our contribution to this strand of research is in documenting interaction across different groups of users at high frequencies.

The third strand of research investigates the role of sentiment in the transmission of information (e.g., Heath, 1996; Chen and Lurie, 2013). In political discourse on Twitter, it is observed that tweets containing a high degree of emotionality reach wider readership and are more likely to be disseminated (Kim and Yoo, 2012; Dang-Xuan et al., 2013; Stieglitz and Dang-Xuan, 2013). There is also evidence for the negativity bias in information consumption that people react faster to the information having emotionally negative tone (e.g., Baumeister et al., 2001). For example, Meffert et al. (2006) find the existence of negativity bias in voters' information selection, processing and recall during a political campaign. However, it is not always the case. During the 2012 U.S. General Election, voters are more exposed to the content mentioning their favorite

candidates with positive sentiment (Halberstam and Knight, 2016). In a contribution to this line of research, we show how sentiment intensity of messages posted by bots influences reactions of “human” users.

This paper also contributes to the fourth strand of literature on the presence of bots in social media. It has been documented that social bots become increasingly sophisticated and can mimic human behaviors, making bots detection more difficult (Ferrara et al., 2016; Haustein et al., 2016). We contribute to this literature by documenting the behavior of bots as well as their influence on humans.

The rest of this paper is organized as follows. In the next section, we describe the dataset and how we collected data. Section 3 presents our empirical strategy and results. Section 4 concludes and provides implications.

## **2. Data**

This section describes how data were collected and filtered and presents a summary of the dataset. We distinguish the tweeting activities of bots and humans to examine later how these two types of agents interact. Next, we perform the sentiment analysis and classify the tone of a tweet into positive, neutral, or negative. Finally, we explore the power of tweeting intensity to predict outcomes of the E.U. Referendum and the 2016 U.S. Presidential Election.

### *2.1. Data collection and cleaning*

The data for analysis were collected using Twitter Streaming API. API can be viewed as a tool for collecting data directly from Twitter in effectively real-time. A user sets selection criteria (e.g. keywords or location) and Twitter sends a sample of selected tweets as they happen. In this study, we make requests to collect tweets that contain certain keywords and leave the connections open to collect as many tweets as possible during the harvested periods. Each retrieved tweet contains the plain text of the tweet as well as information about users like user ID (user name) and other fields such as date, source, location, friend counts, follower counts, URL, retweet counts, etc.

The Brexit-related tweets are collected from 24 May 2016 to 17 August 2016. Tweets are tracked if they contain the keyword “Brexit”. The 2016 U.S. Election-related tweets are collected from 8 October 2016 to 8 December 2016. We collected all tweets containing the following keywords: “Election2016”, “BlackLivesMatter”, “CampaignZero”, “ClintonEmails”,

“ImWithHer”, “NeverClinton”, “FeelTheBern”, “CruzCrew”, “MakeAmericanGreatAgain”, “Trump”, “Clinton”. Given that we collect tweets during high intensity events, our sample is likely to have about 1% of all tweets with keywords mentioned above.

The screening and cleaning process is as follows. First, we process each tweet to extract the relevant content and store in a new tweet content variable. Specifically, we exclude special characters such as link tokens (starting with “http://”, “https://”, “www.”) or user identifier tokens (starting with “@”) from the tweet content. Second, we do not include tweets that contained only links or URLs in our analysis.<sup>3</sup> Third, we separate tweets whose language description is English from those whose language description is other languages. Finally, we adopt the approach proposed by Howard and Kollanyi (2016) and Kollanyi et al. (2016)<sup>4</sup> and define campaign endorsement for each tweet based on the hashtags specified in Appendix Table A1. After screening, our sample contains about 2.7 million tweets for the EU Referendum and 7.2 million tweets for the 2016 U.S. Presidential Election.

In the next step, we identify original tweets (i.e., the tweets were created rather than copied) and their retweets. First, we screen each tweet’s content in the original dataset and create a new indicator variable *RT* which equals 1 if the tweet starts with “RT @” (which means a retweet in the Twitter language; that is, a repost of an original tweet) and 0 otherwise. Next, we extract the content after “@” and before the main text and refer it as *RT\_from*. This *RT\_from* is simply the user name of the Twitter account from which the tweet was retweeted. After these steps, we could identify (1) the original tweets, (2) their direct retweets, and (3) their indirect retweets. An example of this process is provided in Appendix Table A2.

## 2.2. Identification of bots

Previous research (e.g., Chu et al., 2010; Sivanesh et al., 2013; Cook et al., 2014) documents several patterns that help distinguish bots and human users. First, a human agent tends to be more

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<sup>3</sup> This criterion effectively removes Twitter accounts of many standard media (e.g., BBC, Times) because these media typically post only links (URLs) to articles on their Twitter pages.

<sup>4</sup> In these studies, the authors analyse all Brexit-related and U.S. Election-related tweets and use some specific hashtags to define supporting sides. Our approach is different in two ways. First, we do not analyse the tweets that contain only a URL even if the URL includes the hashtags. Second, we do not include some of the hashtags/keywords used by Howard and Kollanyi (2016) and Kollanyi et al. (2016) such as #Brexit, #Trump, or #Clinton to classify tweets as being in favour or against a side of a campaign because these hashtags have been often used to support both sides of the argument.

active in tweeting during the regular workdays and during the daytime while the daily and hourly tweeting activity of a bot agent is even. Second, bots often tweet the same content many times while humans do not. Given the aggressive use of social bots during the political events like election, previous studies also suggest some event-specific criteria to detect bots. For example, bot accounts are more likely to be created just on or about the event announcement date. Further, bot agents could have the inactivity before and after the event while create mass tweets or retweets on event-specific days and times.

Building on these earlier results, we use the following procedure to classify twitter accounts into humans and bots. Consider a given Twitter account on a given day. We flag this account/day as a potential bot if *any* of the following conditions is satisfied.

First, Haustein et al. (2016) document that the average number of daily tweets for a bot is about 5 (standard deviation  $\approx 5$ ) while the daily average for humans is about 2 (standard deviation  $\approx 2$ ). Given these differences in the volume of tweeting, our first criterion is an unusually large number of tweets in a day created by an account. For the Brexit sample, we choose the threshold of 10 which is about 4 standard deviations above normal human activity. Because the Twitter usage including the number of Twitter users and the tweet volumes in the U.S. is substantially greater than that in the U.K., we raise this threshold to 15 for the U.S. Election sample.

Second, since the period from 00:00 to 06:00 is often considered as the inactive time for humans, any actively tweeting activities during this time period could be viewed as “abnormal tweeting time”. We flag an account as a potential bot if the account creates at least five tweets and 10 tweets during the abnormal tweeting time on a given day for the Brexit sample and the U.S. Election sample, respectively. Abnormal time is defined based on British Summer Time for U.K. or Eastern Time for U.S.

Third, previous studies in computer science (e.g., Lee et al., 2010; Chu et al., 2012) suggest that one of bots’ characteristics is to repeatedly post identical messages. Thus, we flag an account as a potential bot if the number of tweets with the same content per day is 3 or more for both U.K. and U.S. samples.

Fourth, Twitter bots may be created and used for particular events. Hence, the next criterion examines newly created accounts. We define an account as “newly created” if the account is created on or after 20 February 2016 for the Brexit sample when the Referendum was announced and 15 July 2016 for the 2016 U.S. Election sample when by-then Republican presumptive

nominee Donald Trump announced his vice-presidential running mate. These accounts are flagged as potential bots if they have an unusually high average daily tweet volume relative to what one might have expected for a twitter account of that age. The chosen thresholds are 10 and 15 for the Brexit and U.S. Presidential Election samples, respectively.

If an account is flagged as a bot for majority of days (that is, more than 50 percent of days) during its lifetime in the sample, then the account is defined as a bot; otherwise the user is defined as a human. We experimented with variations of these criteria (e.g., a user is defined as a bot if tweeting activities are observed for at least three days and on more than 50 percent of days tweeting activities match all four criteria; raising or lowering the threshold point for each criterion) and we found similar results in our regression analysis. To check the validity of our procedure to identify bots, we compare our bot definition with bot detection based on an online classification tool called Botometer (formerly BotorNot).<sup>5</sup> We find that these two approaches have 90% consistency in classification of accounts into bots and non-bots in our U.K. sample.

Since there is a growing concern about the influence of foreign nations like Russia in the 2016 U.S. Presidential Election and Brexit through social media, we also examine the behavior of Twitter accounts that self-declare Russian as their profile language. Obviously, having Russian as the profile language does not imply that an account is sponsored by the Russian government. However, Twitter offers default language based on the language of the operating system of a person who wants to open a Twitter account and thus the location of the person may be unintentionally disclosed.<sup>6</sup> If a large volume of tweets is generated from accounts that have Russian as the profile language, one may suspect that these accounts played a potentially unusual role.

### 2.3. *Dynamics of Twitter posts*

Figure 1 illustrates the evolution in the daily and hourly volumes of Brexit-related tweets for humans (Panels B and D) and bots (Panels A and C). There is a significant increase in the number of tweets created by humans on the 23<sup>rd</sup> and 24<sup>th</sup> June 2016. While bots also show more intensity around these dates but the increase is much more modest. Interestingly, the daily volume of pro-leave tweets was always higher than the daily volume of pro-remain tweets. This gap was greatest

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<sup>5</sup> This tool is developed by researchers from Indiana University and Northeastern University. Botometer tool cannot classify the accounts that have been suspended or deactivated.

<sup>6</sup> A user may manually change the language and location.



during the time around the E.U. Referendum day: between 00:00 and 06:00 on the 24<sup>th</sup> June, the difference in the hourly pro-leave tweets and pro-remain tweets reached its peak of about 10,000 tweets. There is a clear pattern in humans' hourly tweeting volume: human accounts are more active between 6 am and 6 pm and they show considerably smaller intensity in other hours.<sup>7</sup> In contrast, we do not observe any clear pattern in the hour-by-hour tweeting activity of bots.

The time series of 2016 U.S. Election-related tweets are showed in Figure 2. Most of the time, the number of pro-Trump tweets exceeded the daily volume of pro-Clinton tweets. A large increase in pro-Clinton tweets only appeared during the time running up to the Election Day. Specifically, approximately 5 days before and on the Election Day, the number of pro-Clinton tweets soared with the peak of nearly 10,000 tweets per hour and was higher than the number of pro-Trump tweets. Comparing the differences in the number of tweets created by the two sides before and after the voting day, we observe a significant reduction in the hour-by-hour gap between two periods. Note that the intensity of tweeting activity declines sharply after the Election day while in the U.K. the post-referendum decline was more spread out in time.

Figure 3 documents several patterns for Twitter accounts with Russian as the profile language. First, there is a clear difference in the volume of "Russian" tweets between the Brexit sample and the 2016 U.S. Presidential Election sample. For the Brexit sample, the massive volume of "Russian" tweets was created only a few days before the voting day, reached its peak during the voting and result days, then dropped immediately afterwards. In contrast, the volume of "Russian" tweets was relatively high both before and after the U.S. Election Day. In addition, during the running up to the 2016 U.S. Presidential Election, the number of bots' "Russian" tweets dominated the ones created by humans while the difference was modest for other times. Finally, after the 2016 U.S. Presidential Election Day, tweets by "Russian" bots dropped sharply before the new wave of tweets was created.

#### *2.4. Sentiment of the tweets*

Baumeister et al. (2001), Kim and Yoo (2012), Stieglitz and Dang-Xuan (2013) and others show that the intensity of information flows can depend on sentiment (tonality) of messages. To measure

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<sup>7</sup> Appendix Figure B1 shows how intensity of tweeting activity by humans and bots changes by hour of the day and by day of the week.

the intensity of sentiment, we use TextBlob, a publicly available text-processing tool written in Python, to get a polarity score for each tweet. See Loria (2018) for TextBlob details.

TextBlob can perform various tasks such as part-of-speech tagging, noun-phrase extraction, sentiment analysis, spelling correction, text translation and many more. The analysis using PatternAnalyzer in TextBlob returns the polarity score between -1 and 1. A score in  $[-1,0)$  represents negative sentiment, a score in  $(0,1]$  represents positive sentiment, and a score of 0 refers to neutral sentiment. Examples of how TextBlob works in our Brexit sample are in Appendix Table A3.

Overall, the volume of emotionally-colored tweets was relatively moderate: neutral messages are the most prevalent (the average share is 50% and 61% for the Brexit and U.S. Election samples, respectively). Messages with positive sentiment have the second place (the average share is 33% and 25% for the Brexit and U.S. Election samples, respectively). Negative messages are the least frequent (the average share is 17% and 15% for the Brexit and U.S. Election samples, respectively). The distribution of scores is reported in Appendix Figure B2.

Figure 4 shows the daily volume of tweets by sentiment and type of user for the Brexit sample and the 2016 U.S. Presidential Election sample. This pattern is observed for both humans and bots. The daily volumes by sentiment tend to comove strongly. We find similar results when we focus on the hourly volume of tweets around the voting dates and when we split the sample by the sides of the campaign (see Appendix Figures B3-B5). This co-movement suggests that the distribution of sentiment was approximately constant during both campaigns.

### *2.5. Predictive power of public opinions on Twitter*

Previous studies show that Twitter activity may have predictive power for electoral outcomes (e.g., Bermingham and Smeaton, 2011; Tumasjan et al., 2011; Burnap et al., 2016). To explore whether this is the case in our sample, we compare support in Twitter activity and the actual shares of votes received by the sides of the campaigns at the regional level.

To construct the former, we use the location of Twitter users to measure how intensively a given geographical location (a state for the U.S. and a region for the U.K.) supports a given side. Takhteyev et al. (2012) documents that 75 percent of Twitter accounts in their large sample report geographical locations of their owners. We find that a very similar share of users reports their location in our sample. While owners may choose locations different from where they actually

reside, available evidence (e.g., Takhteyev et al., 2012; Haustein and Costas, 2014) suggests that, while imperfect, this information is useful for determining geography of twitter users.

Once the location of users is established, we compute the share of pro-leave “human” tweets in total “human” tweets on the day before Referendum for the Brexit sample and the share of pro-Trump “human” tweets in total “human” tweets on the day before the vote date for the 2016 U.S. Presidential Election sample. Figure 5 shows that these shares are highly correlated with the shares of votes received by the corresponding platform.<sup>8</sup> These results suggest that Twitter activity may be a useful gauge for electoral outcomes in our samples.<sup>9</sup>

### **3. Interactions between bots and humans**

This section examines how information flows across different types of users. The main focus of our analysis is how bots can influence Twitter activity of humans. We use two approaches to measure direction and intensity of the flows. First, we study how frequently a user type retweets (i.e., re-posts) messages of the other user type. Second, we employ time-series tools to investigate how bots’ messages (original and retweets) generate humans’ messages (original and retweets).

#### *3.1. Retweeting*

Similar to other social media, Twitter allows users to repost existing messages. Typically, reposting (retweeting) a message means that a user wants to spread the message in his or her social circle. Messages with many retweets are often labeled as popular/trending and, as a result, have higher ranking/priority in internet searches. In other words, a message with many retweets is often treated as important. Because retweeting a message generates a link from a user who originated the message to a user who reposted it, we can observe the direction of the information flow. Thus, retweeting provides us with measures of intensity and direction for interaction between different types of users.

To understand our approach, consider the following fictitious example. Suppose an original (i.e., not a copy) tweet supporting the leave campaign appears at 1 pm. We compute how many retweets between 1pm and 1:10pm this tweet generated by human accounts and by bot accounts.

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<sup>8</sup> Because voters could write in candidates in the U.S., the actual votes by the U.S. states are calculated using this formula: Actual vote = Votes for Trump/(Votes for Trump + Votes for Clinton).

<sup>9</sup> We find similar results when we weight tweets with sentiment.

Then we count the number of (new) retweets by humans and bots that were generated for the original tweet between 1:10pm and 1:20pm. This procedure is continued at ten-minute intervals for 2 hours after the original tweet appeared. The resulting path provides us with an impulse response function for the original tweet. We repeat this procedure for all original tweet and compute the average impulse response function. Figure 6 reports these average impulse response functions by type of users who generated original tweets and by type of users who retweeted original messages.<sup>10</sup>

Panels A and B of the figure show that, relative to humans, bots are not very active in retweeting. Indeed, the intensity of bots' retweeting is an order of magnitude smaller than retweeting activity of humans. Humans are most active in retweeting in the first 10 minutes right after the time when original tweets are generated. The number of new retweets reduces over time and reaches a stable level within two hours. The results are similar when we restrict the sample to include only original messages generated by humans (Panels C and D) or to include only original messages generated by bots (Panels E and F).<sup>11</sup> Note that humans react much more strongly to tweets generated by other humans than to tweets generated by bots. In contrast, bots are equally passive in retweeting messages of humans and other bots. For instance, during the first 10 minutes since humans post original tweets about Brexit, the number of retweets made by humans is significantly higher than that made by bots (35 retweets vs. 2 retweets for every 10 original tweets, respectively). However, during the period from 110 to 120 minutes since the original tweets were posted, only one retweet is made by other human agents and bots tend to not retweet at all.

These patterns lead us to three tentative conclusions. First, information flows are most intensive between humans, while information flows across bots are weak. Second, information flows from bots to humans are tangible while information flows from humans to bots are very weak. Third, reactions tend to be relatively short-lived in the sense that the vast majority of the reaction is completed within two hours.

To further investigate how humans reacted to the messages created by bots, we restrict our samples to bots' original tweets and check the daily retweets made by bots and humans (dashed line in Figure 7). We find that for a typical bots' tweet that was retweeted, about 80-90 percent of

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<sup>10</sup> In the Brexit sample, the average number of retweets per original tweet is 0.5 and the standard deviation is around 4. In the U.S. Election sample, these figures are 0.3 and 12, respectively. Because most tweets generate few or no retweets, we restrict the sample to relatively popular original tweets (the ones get more than 5 retweets) to have meaningful variation over time.

<sup>11</sup> Similar patterns are also observed when we separate tweets by sentiment.

retweets were made by humans. However, when we restrict our samples to humans' original tweets and check the daily retweets made by bots and humans (solid line in Figure 7), we find that bots account for only 5 to 10 percent of retweets generated in response to humans' original tweets. Thus, humans tended to spread messages that were originally generated by bots while bots did not spread messages generated by humans.

These results are consistent with the view that humans had little (if any) effect on bots while bots had a perhaps limited, but tangible effect on humans. The asymmetric response of humans to posts consistent and inconsistent with their views suggests that social media can create “echo chambers” fostering amplification of messages within a group of similarly-minded people and inhibiting communication of people with different views. These results also suggest that bots are not likely to be successful in persistently moving tweeting activity of humans. The short duration of the response is consistent with the fast news cycles in social media (e.g., Kwak et al., 2010; Yoo et al., 2016) and/or low information rigidity (Coibion and Gorodnichenko, 2012).

### *3.2. Time series analysis*

While the analysis in the previous section is informative, it is focused on reposts of original messages. Obviously, interaction between different types of users may also happen via generation of new messages. In this subsection, we use standard tools of time-series analysis to construct impulse responses of all messages (that is, retweets and new posts) by a type of users to a message generated by a given type of users.

This exercise relies on two key ingredients. First, we build on our earlier findings for retweets and assume that humans can respond contemporaneously to bots while bots do not respond contemporaneously to humans. Second, to ensure that this identifying assumption holds, we use data at 10-minute intervals. Apart from strengthening our identification, this short duration of time intervals allows us to control for low-frequency variation in the volume of tweeting activity (e.g., days closer to the vote date have higher volume than more distant days).

#### A. Econometric specification

To estimate impulse responses flexibly, we use local projections method developed by Jordà (2005). To see how this method works, suppose we are interested in estimating reactions of humans

supporting campaign  $X$  to bots advocating campaign  $X$  and to bots advocating campaign  $Y$ . The method amounts to estimating  $h = 0, \dots, H$  regressions of the following type:

$$\begin{aligned} \ln Human_{t+h,d}^{(X)} = & \sum_{k=0}^K \alpha_{X,k}^{(h)} \ln Bot_{t-k,d'}^{(X)} + \sum_{k=0}^K \beta_{X,k}^{(h)} \ln Bot_{t-k,d'}^{(Y)} \\ & + \sum_{k=1}^K \gamma_{X,k}^{(h)} \ln Human_{t-k,d'}^{(X)} + \sum_{k=1}^K \phi_{X,k}^{(h)} \ln Human_{t-k,d'}^{(Y)} \\ & + \psi_{X,d}^{(h)} + Seasonal_{X,td}^{(h)} + error_{X,td}^{(h)} \end{aligned} \quad (1)$$

where  $t$  and  $h$  index ten-minute intervals,  $d$  indexes the day of a campaign,  $Human_{t+h,d}^{(X)}$  is the volume of new tweets generated by humans supporting campaign  $X$  during the  $t + h$  ten-minute interval on day  $d$ ,  $Bot_{t-k,d'}^{(C)}$  is the volume of new tweets by bots supporting campaign  $C = \{X, Y\}$  during the  $t - k$  ten-minute interval on day  $d'$  where  $d' = d$  if the  $t - k$  interval is on the same day with  $t$  and  $d' = d - 1$  if the  $t - k$  interval is on the day proceeding  $t$ . Because there is considerable variation in tweeting activity during the day and week, we include  $Seasonal_{td}^{(h)}$ , a set of “seasonal” dummy variables. Specifically, for each 1-hour interval during a 24-hour day period we have a dummy variable; note that each weekday (Monday, Tuesday, etc.) is allowed to have a potentially different 24-hour profile of intra-day activity. Finally,  $\psi_d^{(h)}$  is a dummy variable equal to one if the day of campaign is equal to  $d = \{-30, -29, \dots, 0, \dots, 29, 30\}$ .

Note that in this specification, the lag polynomial of humans supporting campaign  $Y$  ( $\ln Human_{t-k,d'}^{(Y)}$ ) starts with  $k = 1$  while the lag polynomials for bots start at  $k = 0$ . This timing means that we allow bots to have a contemporaneous effect on humans and bots do not respond to humans. Consistent with earlier studies using the Jordà approach, we use Newey-West standard errors to account for serial correlation of the error term for  $h \geq 1$ . We use  $K = 24$  for the reported results but our findings are largely unchanged for alternative values of  $K$ .

We compute the impulse response to bots supporting campaign  $X$  as  $\{\hat{\alpha}_{X,0}^{(h)}\}_{h=0}^H$  and the impulse response to bots supporting campaign  $Y$  as  $\{\hat{\beta}_{X,0}^{(h)}\}_{h=0}^H$ . Note that we use logs in specification (1) to transform the volume of tweeting activity (this helps to make the distribution of tweet volume better behaved) so that  $\alpha_0$  and  $\beta_0$  are elasticities. To convert these elasticities into “multipliers” (that is, a tweet from bot  $X$  generates  $N$  tweets by humans supporting  $X$ ), we multiply

$\alpha$  by the ratio  $\overline{Human^{(X)}/Bot^{(X)}} \approx (\overline{Human^{(X)}/Bot^{(X)}})$ , that is, the time-series average of the  $Human_d^{(X)}/Bot_d^{(X)}$  ratio. Correspondingly, the multiplier from bot  $Y$  to human  $X$  is the product of  $\beta$  and  $\overline{Human^{(X)}/Bot^{(Y)}}$ .

One can think of the Jordà method as constructing a moving average representation of a series: the lag polynomial terms control for initial conditions while  $\{\hat{\alpha}_{X,0}^{(h)}\}_{h=0}^H$  and  $\{\hat{\beta}_{X,0}^{(h)}\}_{h=0}^H$  describe the behavior of the system in response to a structural, serially uncorrelated shock. Indeed, if we abstract from variation in initial conditions at time  $t$ , we effectively regress a variable of interest at time  $t + h$  on a shock in a given regime at time  $t$  and thus we obtain an average response of the variable of interest  $h$  periods after the shock, which is precisely the definition of an impulse response.

As discussed in Auerbach and Gorodnichenko (2012), this approach has several advantages over vector autoregressions (VARs). First, it obviates the need to estimate the equations for dependent variables other than the variable of interest and thus we can significantly economize on the number of estimated parameters. Second, it does not constrain the shape of the impulse responses. Third, one can easily test joint hypotheses about paths of estimated impulse response. Finally, specification (1) may be straightforwardly extended in various dimensions to allow for a larger set of controls or for more flexible (potentially non-linear) responses. For example, we are interested in comparing the strength of a reaction to human posts supporting campaign  $Y$  to the strength of a reaction to bot posts supporting campaign  $Y$ . To obtain this comparison, we can estimate

$$\begin{aligned} \ln Human_{t+h,d}^{(X)} = & \sum_{k=0}^K \alpha_{X,k}^{(h)} \ln Bot_{t-k,d'}^{(X)} + \sum_{k=0}^K \beta_{X,k}^{(h)} \ln Bot_{t-k,d'}^{(Y)} \\ & + \sum_{k=1}^K \gamma_{X,k}^{(h)} \ln Human_{t-k,d'}^{(X)} + \sum_{k=0}^K \phi_{X,k}^{(h)} \ln Human_{t-k,d'}^{(Y)} \\ & + \psi_{X,d}^{(h)} + Seasonal_{X,t,d}^{(h)} + error_{X,t,d}^{(h)} \end{aligned} \quad (1')$$

where now the lag polynomial for  $\ln Human_{t-k,d'}^{(Y)}$  starts at  $k = 0$  rather than  $k = 1$  and one can use  $\{\hat{\phi}_{X,0}^{(h)}\}_{h=0}^H$  as the impulse response of humans supporting  $X$  to humans supporting  $Y$  and corresponding measure multipliers are the product of  $\phi$  and  $\overline{Human^{(X)}/Human^{(Y)}}$ . Note that this specification is equivalent to ordering  $Human^{(Y)}$  before  $Human^{(X)}$  in a VAR.

## B. Baseline results

Figure 8 reports estimated impulse responses for the Brexit sample. Panel A of the figure shows the reaction of humans supporting the leave campaign to messages generated by bots supporting the leave campaign and by bots supporting the remain campaign. The response to “remain” bots is generally small with a weak reaction on impact and a modest, positive multiplier in subsequent periods. In contrast, the contemporaneous reaction of “leave” humans to “leave” bots is strong: the multiplier is close to 2, that is, a new bot post generates two new human posts. However, this elevated tweeting activity of humans is short-lived: after approximately 2-4 hours of the bot post we observe little difference in the response of “leave” humans to “leave” bots and to “remain” bots. These patterns are similar to the behavior of humans in retweeting posts thus providing additional evidence of “echo chambers” as well as fast news cycles and/or low information rigidity.

Panel B of the figure plots the responses of “remain” humans to “leave” bots and “remain” bots. Similar to what we observe in Panel A, the reaction of “remain” humans to bots advocating the other side of the campaign is rather mute (the multiplier is close to zero), while the reaction to bots from the same side of the campaign is stronger (the multiplier is about 0.7 which is smaller in absolute terms than the contemporaneous multiplier for “leave” humans in response to “leave” bots). Likewise, the effect of “remain” bots on “remain” humans is rather transitory.

The patterns are similar for the 2016 U.S. Presidential Election (Figure 9). Human supporters of the Trump (Clinton) campaign are more reactive to messages posted by bots supporting the Trump (Clinton) campaign than to messages posted by bots supporting the Clinton (Trump) campaign. In a similar spirit, the reactions are not persistent and most of the response happens within a few hours after a message appears.

## C. Sentiment

As discussed above, the intensity of human responses may vary with the strength of the sentiment in messages posted by bots. To study this possible heterogeneity in responses, we modify specification (1) as follows:

$$\begin{aligned} \ln Human_{t+h,d}^{(\tau,X)} &= \sum_{s \in \{positive, neutral, negative\}} \sum_{k=0}^K \alpha_{X,s \rightarrow \tau,k}^{(h)} \ln Bot_{t-k,d'}^{(s,X)} \\ &+ \sum_{s \in \{positive, neutral, negative\}} \sum_{k=0}^K \beta_{X,s \rightarrow \tau,k}^{(h)} \ln Bot_{t-k,d'}^{(s,Y)} \\ &+ \sum_{s \in \{positive, neutral, negative\}} \sum_{k=1}^K \gamma_{X,s \rightarrow \tau,k}^{(h)} \ln Human_{t-k,d'}^{(s,X)} \end{aligned}$$



$$\begin{aligned}
& + \sum_{s \in \{positive, neutral, negative\}} \sum_{k=1}^K \phi_{X,s \rightarrow \tau, k}^{(h)} \ln Human_{t-k, d}^{(s, Y)}, \\
& + \psi_{X, \tau, d}^{(h)} + Seasonal_{X, \tau, td}^{(h)} + error_{X, \tau, td}^{(h)}
\end{aligned} \tag{2}$$

where  $\tau$  and  $s$  index sentiment (tonality) from positive to neutral to negative. One can now interpret slopes as the strength of the response that may vary by tonality of messages posted by humans and bots. For example,  $\alpha_{X, s \rightarrow \tau, k}^{(h)}$  measures the elasticity of the  $h$ -horizon response of humans supporting campaign  $X$  with sentiment  $\tau$  to bot posts supporting campaign  $X$  with sentiment  $s$ . Then we use appropriate ratios of the regressand to a regressor to convert the estimated elasticities into multipliers. Figure 10 and Figure 11 plot the estimated impulse responses (measured as multipliers) for the U.K. and U.S. samples respectively. By and large, we observe results similar to the baseline results: humans supporting a given side of a campaign tend to react stronger to posts generated by bots supporting the same side and the sentiment of human responses mimics the sentiment of bot posts.

#### D. Humans vs. bots

Our analysis so far has presented evidence consistent with the view that the Twitter activity of bots can affect the Twitter activity of humans who share beliefs advocated by bots. However, bots appear to have a weak effect on humans who have beliefs opposite to what is advocated by bots. Would humans be more effective in reaching across the aisle?

To answer this question, we use specification (1') to compare response multipliers for humans who support a given campaign side to posts by humans and bots from the other camp. Panel A of Figure 12 shows response multipliers of “leave” humans in the Brexit sample to tweets posted by “remain” humans and “remain” bots. We observe that, if anything, bots appear to have larger multipliers than humans. Likewise, “remain” humans appear to have larger multipliers in response to “leave” bots than to “leave” humans (Panel B). The results are similar for the U.S. sample (Figure 13): bots appear to be as effective (if not more effective) as humans in moving humans with opposite views.

These results suggest that while human touch and personal connections may be important, in the world of social media bots and other “strangers” can play an equally important role in bringing together or distancing people with different beliefs. Given that bot traffic is considerably cheaper

than traffic generated by humans (Forelle et al., 2015), one may anticipate ever greater use of bots in political campaigns as well as various attempts of humans to shield themselves from bots.

#### E. Historical contribution of bots

Our analysis suggests that bots may have tangible influence on the tweeting activity of humans. To quantify the contribution of bots' traffic to the volume of human tweets, we use the method developed by Coibion et al. (2017) to construct historical decomposition of the time series for human tweet volumes. In particular, we are interested in constructing counterfactual time series of human tweets that would have been observed in the absence of bot traffic.

To implement the Coibion et al. method, we make two departures from specification (1). First, we use innovations to bots' volume of tweets rather than the level of bots' tweets on the right-hand side of specification (1).<sup>12</sup> This step is necessary to have the dynamics of human tweet activity as a moving average process in terms of "bot" shocks so that we can construct a counterfactual dynamic of human tweet activity when we turn off "bot" shocks. We continue using lags of human tweets in levels (rather than in shocks). As a result, we have a combination of moving average (MA) terms (that is, current and lagged values of bot shocks) and autoregressive (AR) terms (that is, lags of human traffic supporting campaign  $X$  and lags of human traffic supporting campaign  $Y$ ). To ensure that we have enough persistence in this vector ARMA representation of the stochastic process for human tweets, we increase the number of lags  $K$  from 24 to 99. We find that impulse responses based on this modification are virtually identical to the impulse responses based on specification (1). Thus, for all practical purposes, this modification does not alter our previous conclusions.

Second, we do not include dummies  $\psi_a$  for each day of a campaign. While these dummy variables are helpful to control for trends in the data, they also demean bots' traffic so that the contribution of bots to the *daily* volume of human tweet activity is zero by construction. Fortunately, we find that removing  $\psi_a$  makes little difference for the estimated impulse response and, therefore, our previous conclusions continue to apply.

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<sup>12</sup> The shocks to bots are constructed as follows. We use specification (1) with bot traffic as the dependent variable and all lag polynomials starting with  $k = 1$ . We estimate this specification separately for bots supporting campaign  $X$  and for bots supporting campaign  $Y$ . The residual of each regression is interpreted as the shock to bots supporting the corresponding campaign.

Note that the dynamics of human tweets are now modelled as a system of two equations with two endogenous variables (e.g., “leave” human tweets and “remain” human tweets) driven by bot shocks (e.g., “leave” bot shocks and “remain” bot shocks) and by shocks to human tweet activity (e.g.,  $error_{leave,td}$  and  $error_{reman,td}$ ). By plugging these shocks into the estimated vector ARMA process, we recover actual time series of human tweet activity. By setting bot shocks to zero, we construct counterfactual dynamics for human tweet activity when bots are not present. To make time series easier to interpret, we aggregate the generated volumes from 10-minute frequency to daily frequency.

Panel A of Figure 14 plots times series of actual and counterfactual (“no bot”) daily volume for pro-“leave” human tweets in the Brexit sample. The difference between the lines is the bots’ contribution. The dynamics of the series suggest that bots had a considerable quantitative effect on the volume of human tweeting with the largest contribution around the vote date. Panel B of the figure documents a similar pattern for pro-“remain” human tweets. Importantly, the two surges generated by bots roughly offset each other: the share of pro-“leave” human tweets in all tweets is similar for actual tweeting volume and counterfactual tweeting volume (Panel C). Specifically, the actually observed share of pro-“leave” human tweets on the day before the vote day is 62.76 percent, while the counterfactual share is 60.69 percent. This is a small absolute difference but one should bear in mind that the Brexit outcome was decided by a small margin (the share of “leave” votes was at 51.9 percent). Our analysis in Section 2.5 indicates that a percentage point increase in the share of pro-“leave” tweets in total tweets is associated with a 0.85 percentage point increase in the share of actual pro-“leave” votes. Hence, the difference between actual and counterfactual traffic could translate into 1.76 percentage points of actual pro-“leave” vote share. Thus, while bots nearly offset each other, the difference could have been sufficiently large to influence the outcome given how close the actual vote was.

Panels D-F replicate the analysis for the 2016 U.S. Presidential Election sample. Similar to the Brexit sample, bots appear to have a considerable contribution to the human tweet traffic. In a similar spirit, the pro-Trump and pro-Clinton human traffic generated by bots nearly offsets each other: actual and counterfactual shares of pro-Trump human tweet volume in total tweet volume are 48.42 and 53.89 percent respectively on the day before the vote date (note that pro-Clinton traffic surged days before the vote date while pro-Trump traffic was more stable so that bots “helped” the Clinton campaign). But again, even this small difference could have played an

important role in the outcome of these close-call elections. Specifically, our analysis in Section 2.5 suggests that a percentage point increase in the share of pro-Trump tweets in total tweets is associated with a 0.59 percentage point increase in the share of actual pro-Trump votes. Therefore, the observed difference between actual and counterfactual pro-Trump tweet shares suggests that 3.23 percentage points of the actual vote could be rationalized with the influence of bots.

#### **4. Concluding remarks**

Social media are a powerful tool for spreading news and information. However, social media might also propagate misinformation and fake news, especially during high-impact events. It is necessary to understand how information is diffused and acquired in social networks as it might affect individuals' decision-making in real life. Furthermore, the rise of bots (automated agents) in social media potentially creates greater risks of manipulation as humans cannot detect bots and thus could be driven and possibly deceived by bots.

This study explores the diffusion of information on Twitter during two high-impact political events in the U.K. (2016 E.U. Referendum, "Brexit") and the U.S. (2016 Presidential Election). Specifically, we empirically examine how information flows during these two events and how individuals' actions might be influenced by different types of agents. We have two key results. First, information about the Brexit Referendum and the U.S. Election is disseminated quickly on Twitter. During these two highly covered campaigns, reaction to new messages is largely complete within 1-2 hours which is consistent with fast news cycles and/or low information rigidity in social media.

Second, we find that individuals are more active in interacting with similar-minded Twitter users. That is, pro-leave users react faster and stronger to the messages created by other pro-"leave" users. We document that the extent to which bots can affect humans' tweeting activities depends on whether bots' information is consistent with humans' preferences. For example, a message by a pro-leave bot generates a response of pro-leave humans and approximately no response of pro-remain humans. Furthermore, bots' messages with a given sentiment largely generate human messages with the same sentiment. These results lend support to the "echo chambers" view that Twitter and other social media create networks for individuals sharing the similar political beliefs so that they tend to interact with others from the same communities and thus their beliefs are reinforced. By contrast, information from outsiders is more likely to be ignored. Consequently,

ideological polarization in social media like Twitter is likely amplified rather than attenuated, which makes reaching consensus on important public issues more difficult.

Since Twitter and other platforms of social media may create a sense of public consensus or support, social media could indeed affect public opinions in new ways. Specifically, social bots could spread and amplify (mis)information thus influencing what humans think about a given issue and likely reinforcing humans' beliefs. Not surprising, bots were used during the two campaigns we study to energize voters and, according to our simple calculations, bots could marginally contribute to the outcomes of the Brexit and the 2016 U.S. Presidential Election.

These two campaigns and subsequent debates about the role of bots in shaping the campaigns raise a number of questions about whether policymakers should consider mechanisms to prevent abuse of bots in the future. Obviously, regulating information flows is an extremely delicate business in a democratic society characterized by diverse views and tolerance for this diversity. However, cherishing diversity does not mean that one should allow dumping lies and manipulations to the extent that the public cannot make a well-informed decision. Where one should draw the line (e.g., improve media literacy, introduce "code of practice" for social networks) is a central question for the society.

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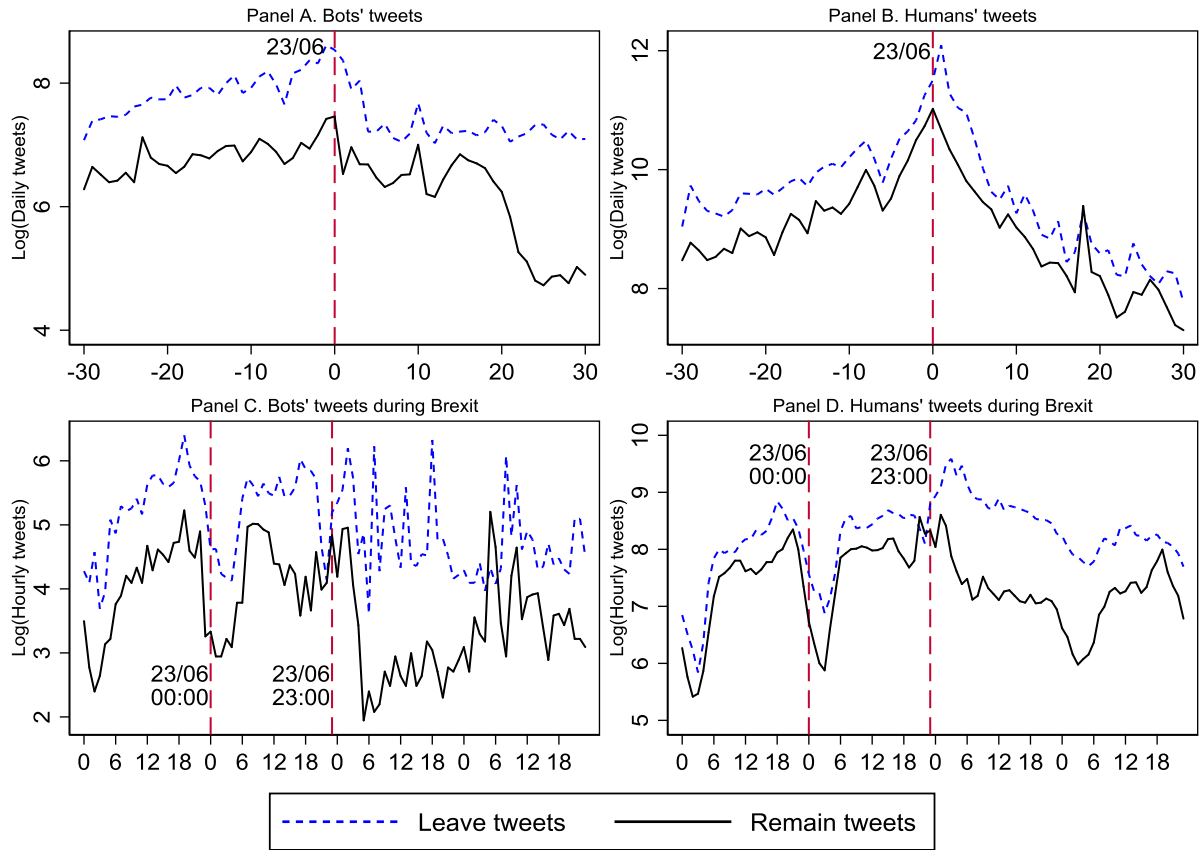
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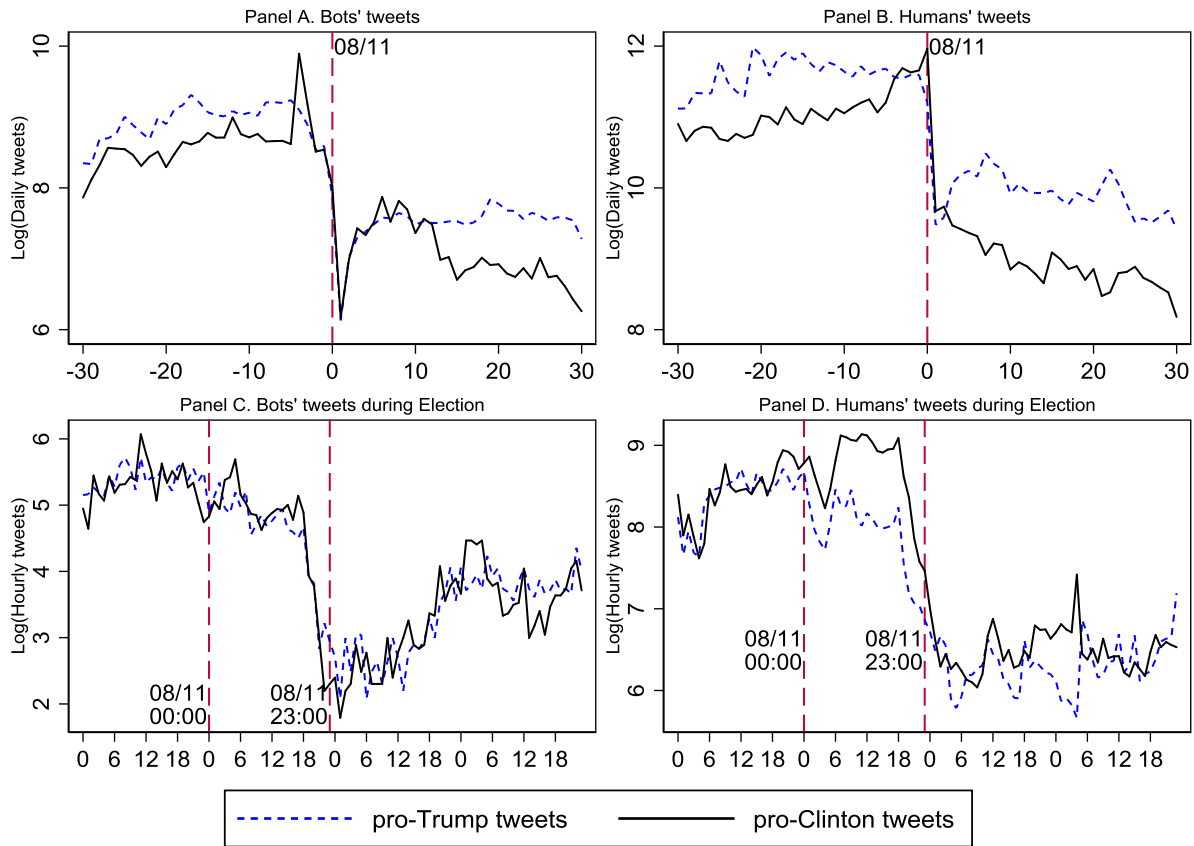
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Figure 1. Dynamics of tweets generated by humans and bots: U.K. Brexit



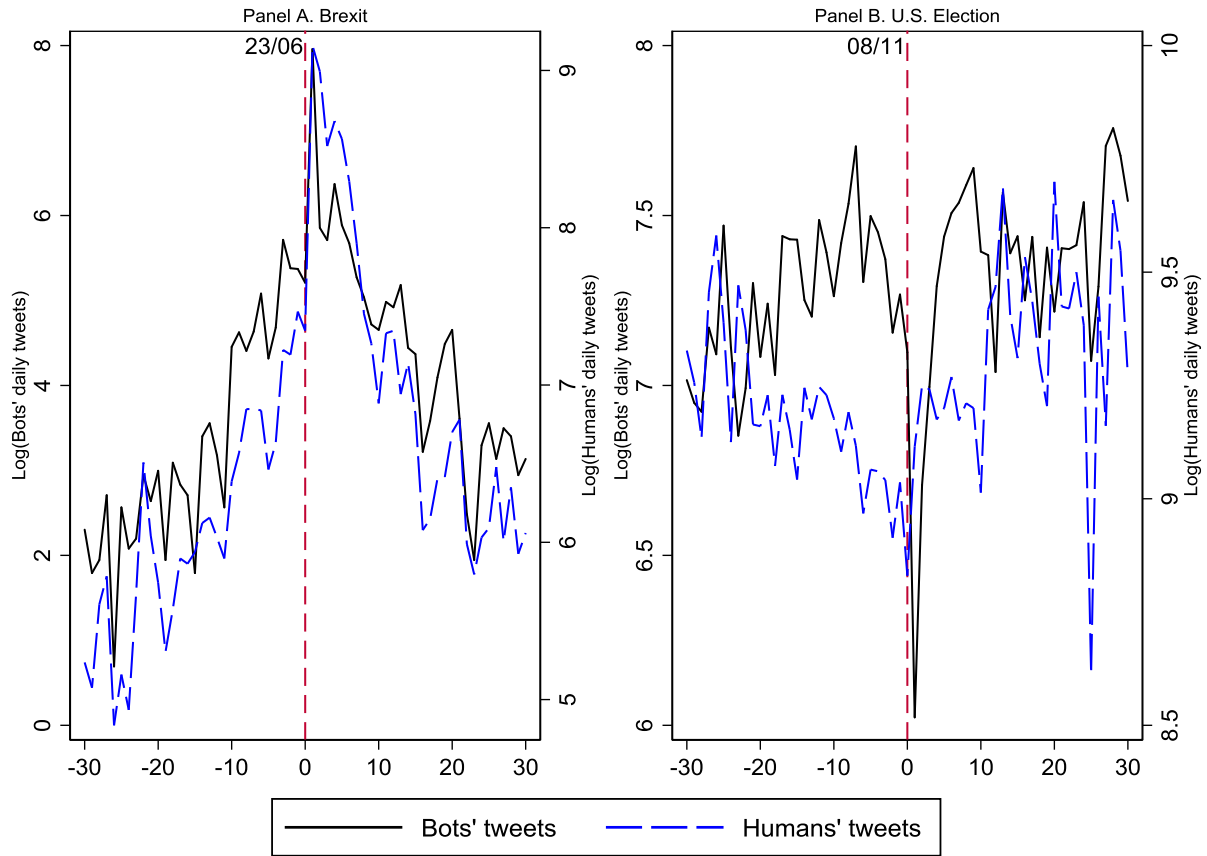
Notes: This figure shows the dynamics of tweets created by bots (Panels A and C) and humans (Panels B and D) for the Brexit sample. Time (horizontal axis) in Panels A and B shows 30 days before and after the Referendum day while time (horizontal axis) in Panels C and D presents hours of days around the event. The dashed blue line and the solid black line show the volumes of pro-leave and pro-remain tweets, respectively.

Figure 2. Dynamics of tweets generated by humans and bots: 2016 U.S. Presidential Election



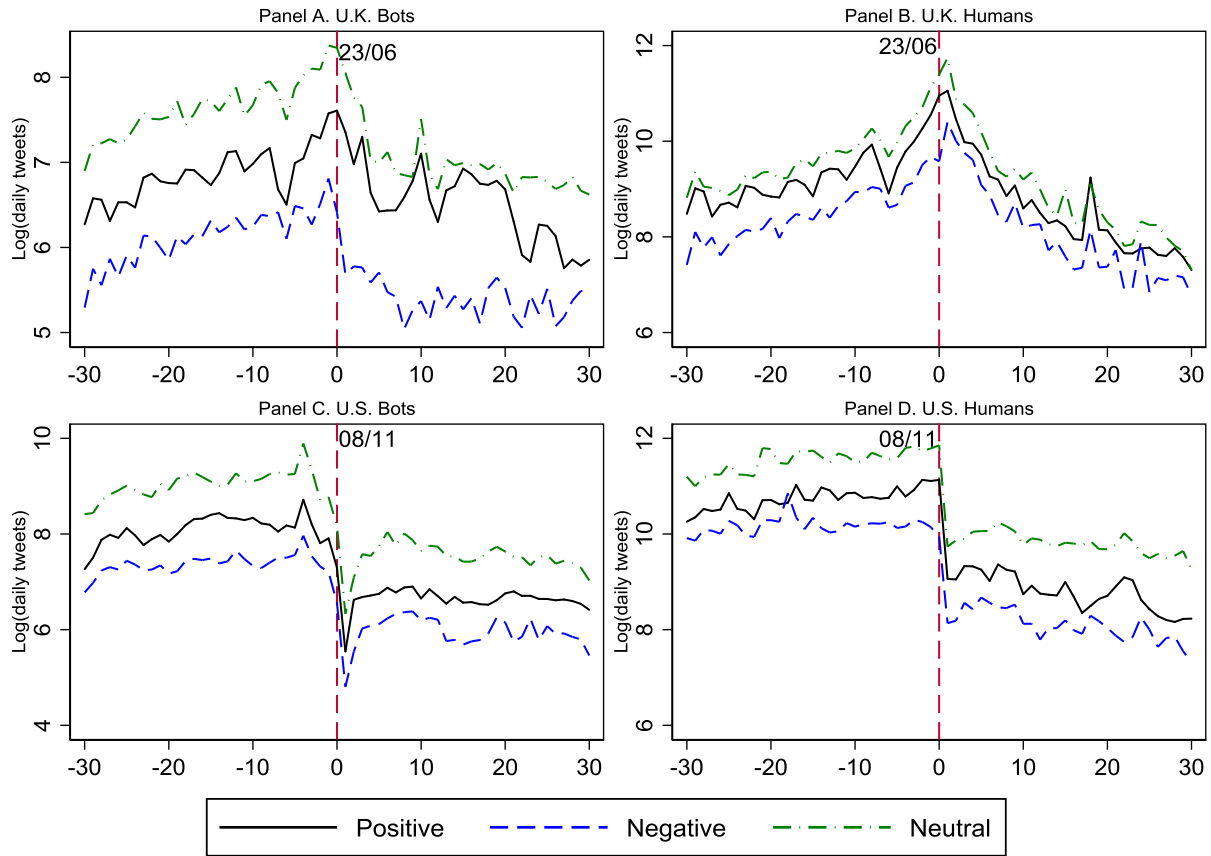
Notes: This figure shows the dynamics of tweets created by bots (Panels A and C) and humans (Panels B and D) for the U.S. Election sample. Time (horizontal axis) in Panels A and B shows 30 days before and after the Election Day while time (horizontal axis) in Panels C and D presents hours of days around the event. The dashed blue line and the solid black line show the volumes of pro-Trump and pro-Clinton tweets, respectively.

Figure 3. Time series of tweets created by Russian-language accounts



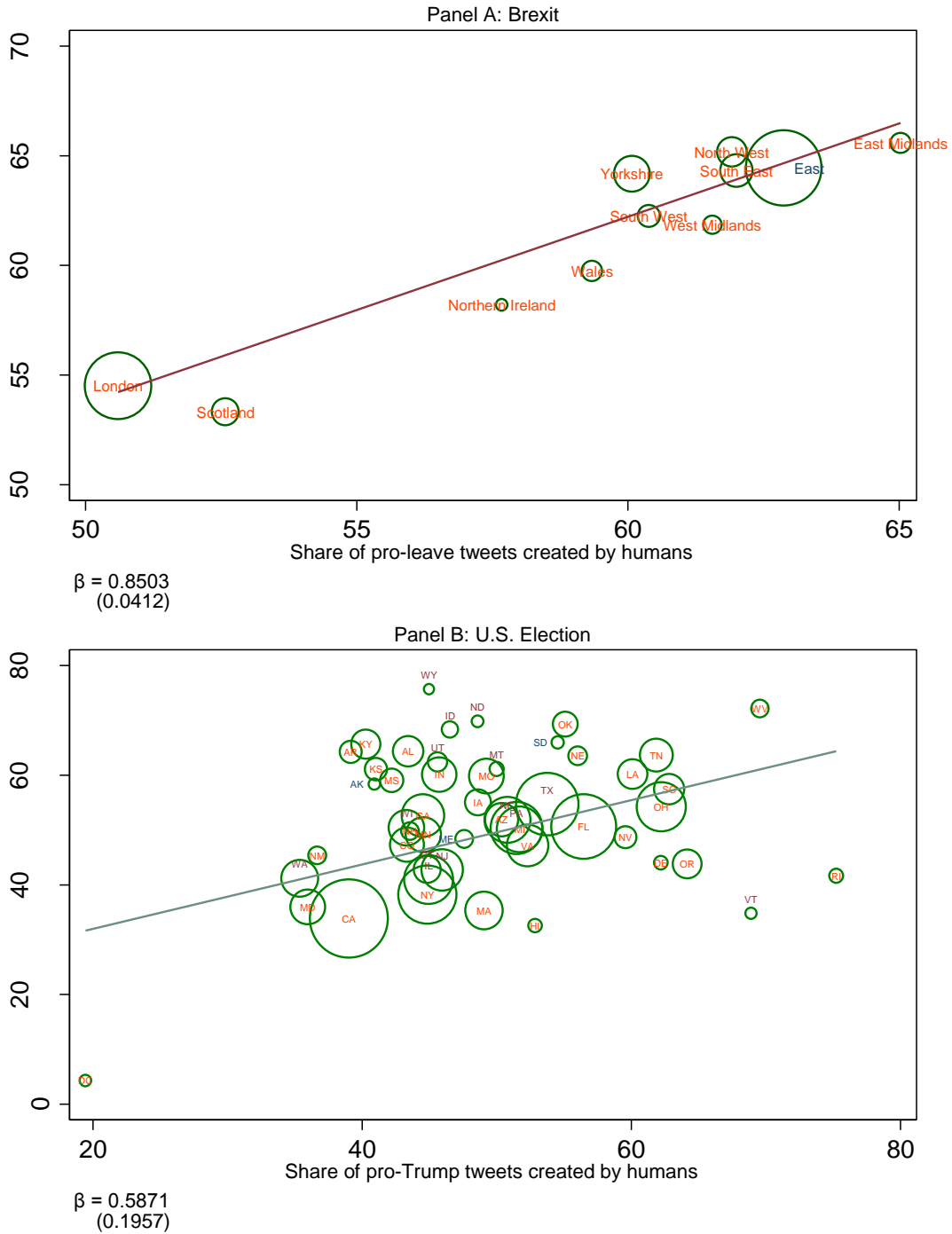
Notes: This figure shows the dynamics of daily tweets created by bot and human accounts that have Russian as the profile language. Panels A and B show the number of daily tweets in logarithm for the Brexit and U.S. Election samples, respectively. The left vertical axis shows the logarithm of the number of bots' tweets while the right vertical axis shows the logarithm of the number of humans' tweets. Time (horizontal axis) shows 30 days before and after the event day. The dashed blue line and the solid black line show the volumes of humans' and bots' tweets, respectively.

Figure 4. Sentiment



Notes: This figure shows the dynamics of tweets with different sentiment created by bots and humans for the samples of Brexit (Panels A and B) and U.S. Election (Panels C and D). Time (horizontal axis) shows 30 days before and after the event day. The dashed blue line, the solid black line, and the dashed-dotted green line show the volumes of tweets with negative, positive, and neutral sentiment, respectively.

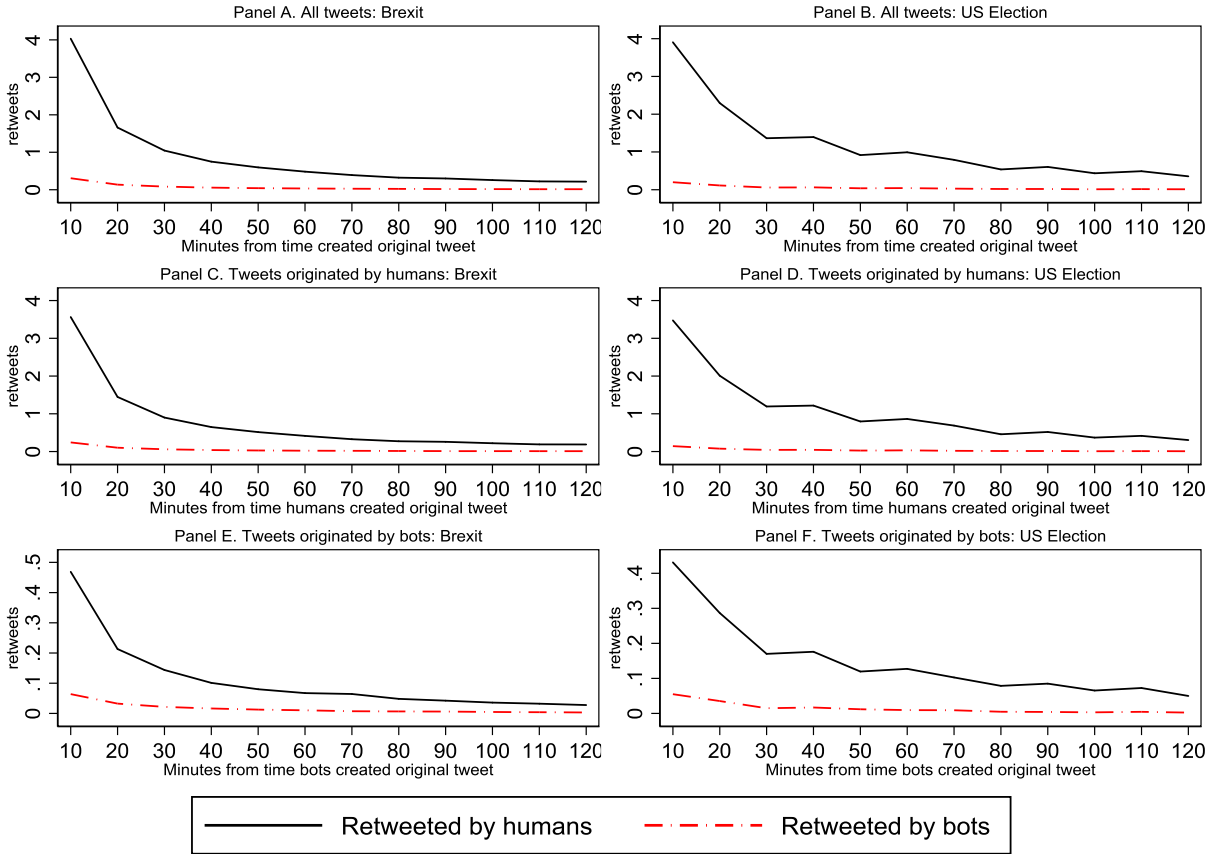
Figure 5. Twitter activity and vote outcomes by geography



Notes: This figure shows the correlation between the shares of humans' pro-leave tweets and the actual vote shares by regions for the Brexit sample (Panel A) and the correlation between the shares of humans' pro-Trump tweets and the actual vote shares for the U.S. Election sample (Panel B).  $\beta$  below each panel shows the estimated slope and standard error (in parentheses).



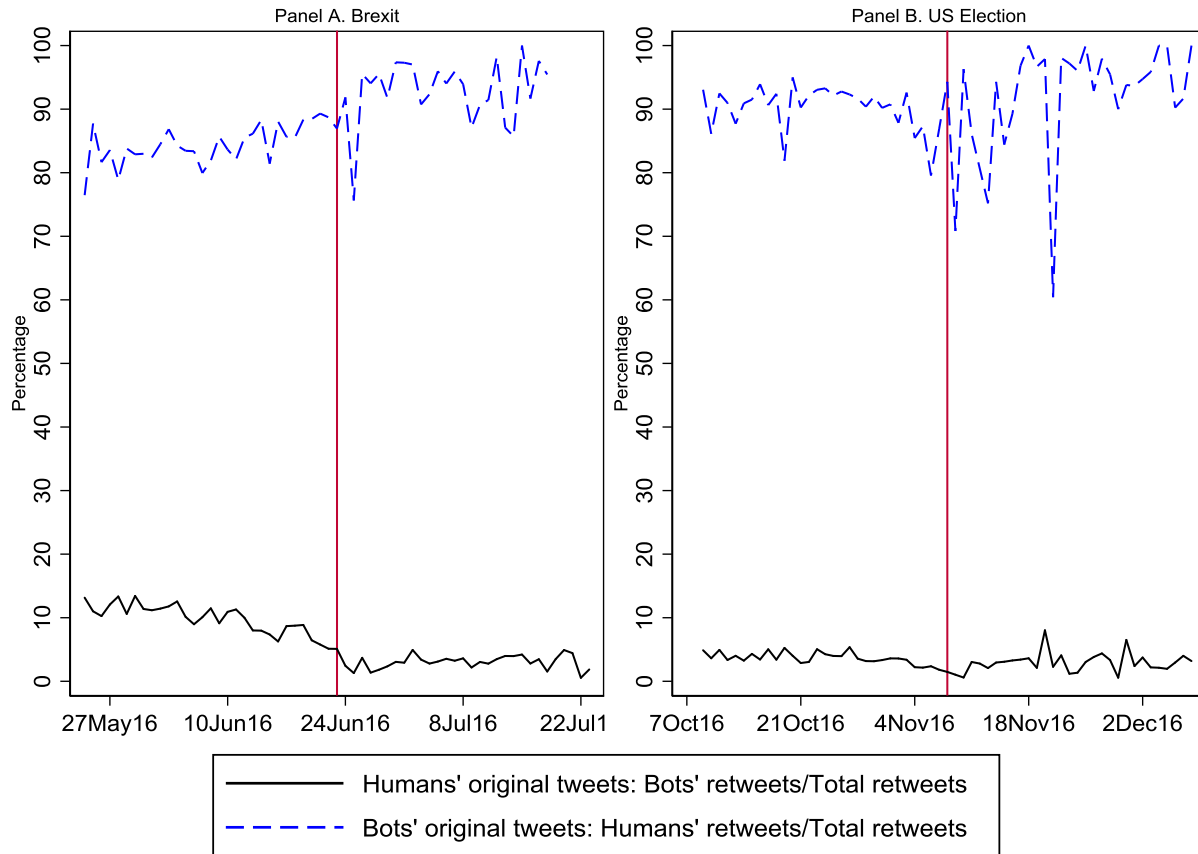
Figure 6. Retweeting activity by humans and bots



Notes: This figure shows the average number of retweets made after 10-120 minutes since the original tweets were created for the Brexit sample (Panels A, C, E) and the U.S. Election sample (Panels B, D, F). Panels A and B show the average retweets of all tweets. Panels C and D show the average retweets of the tweets originated by humans. Panels E and F show the average retweets of the tweets originated by bots. The solid black line refers to the retweeting activities of humans while the dashed red line refers to the retweeting activities of bots. Time (horizontal axis) is in 10-minute intervals.

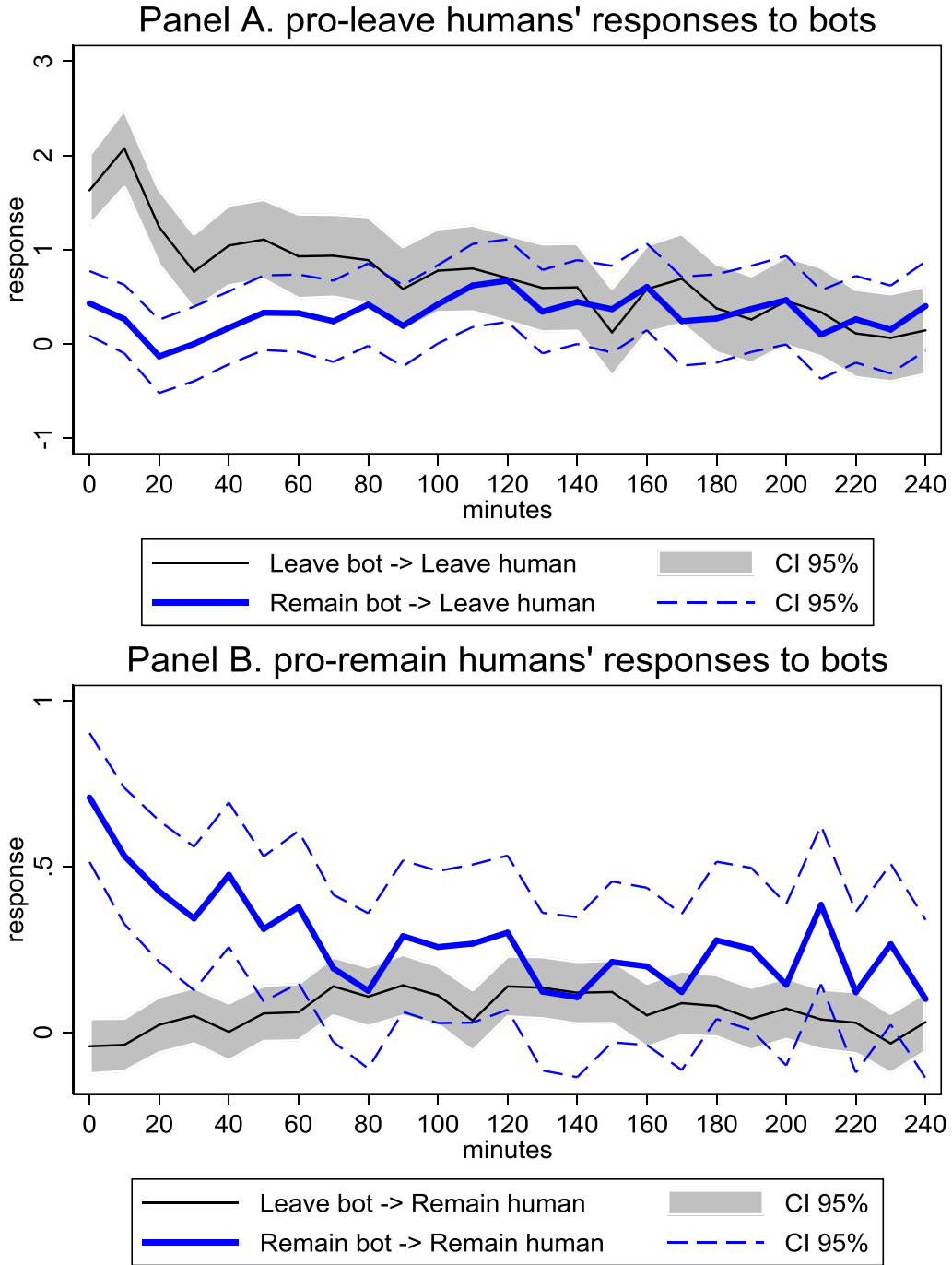


Figure 7. Share of human- and bot-generate retweets by humans and bots



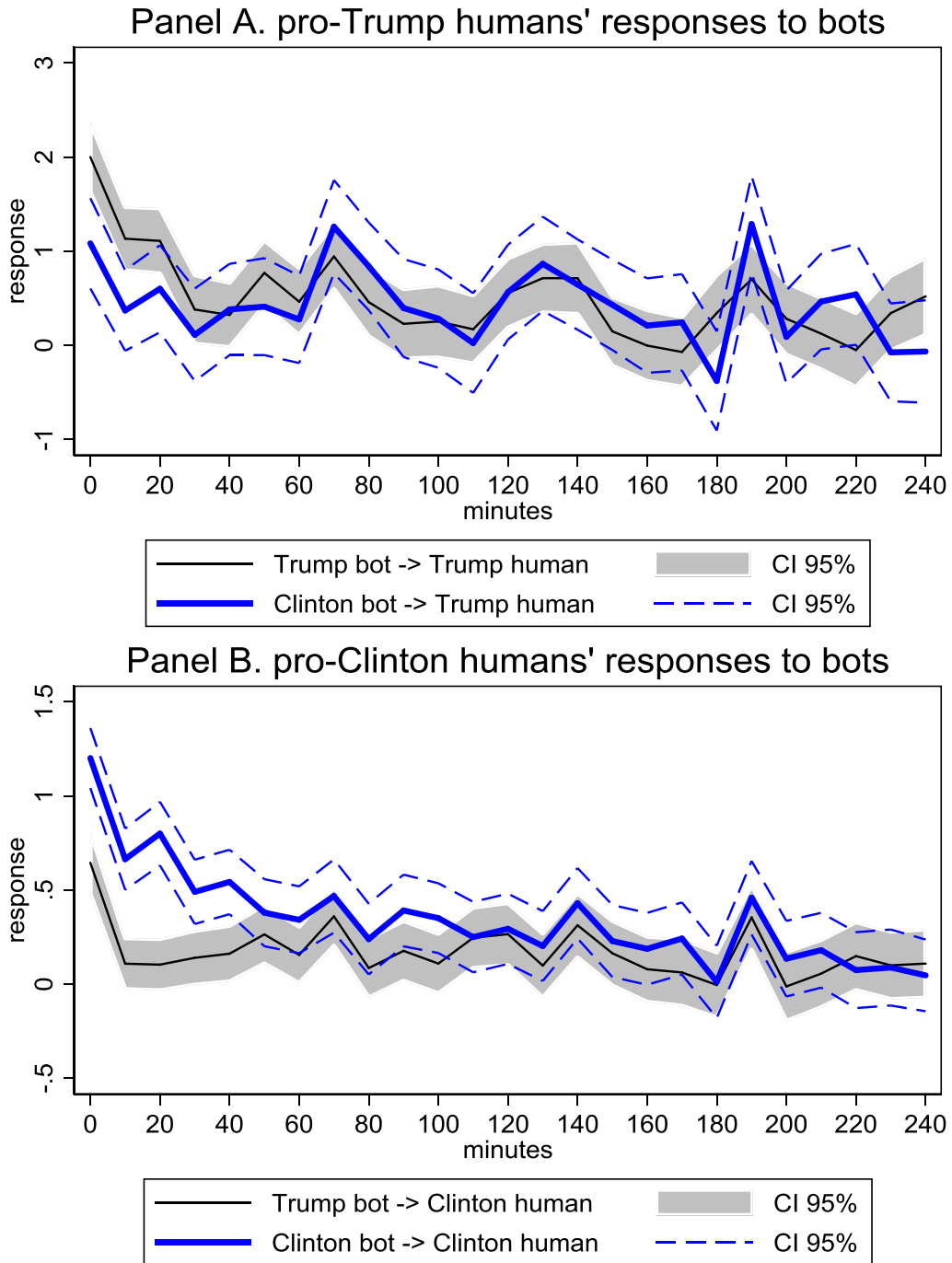
Notes: This figure shows the interactions between humans and bots in retweeting for the Brexit sample (Panel A) and the U.S. Election sample (Panel B). The solid black line refers to the share of the number of retweets made by bots in the case the original tweets were originated by humans. The dashed blue line refers to the share of the number of retweets made by humans in the case the original tweets were originated by humans. Time (horizontal axis) is in days.

Figure 8. U.K. (baseline)



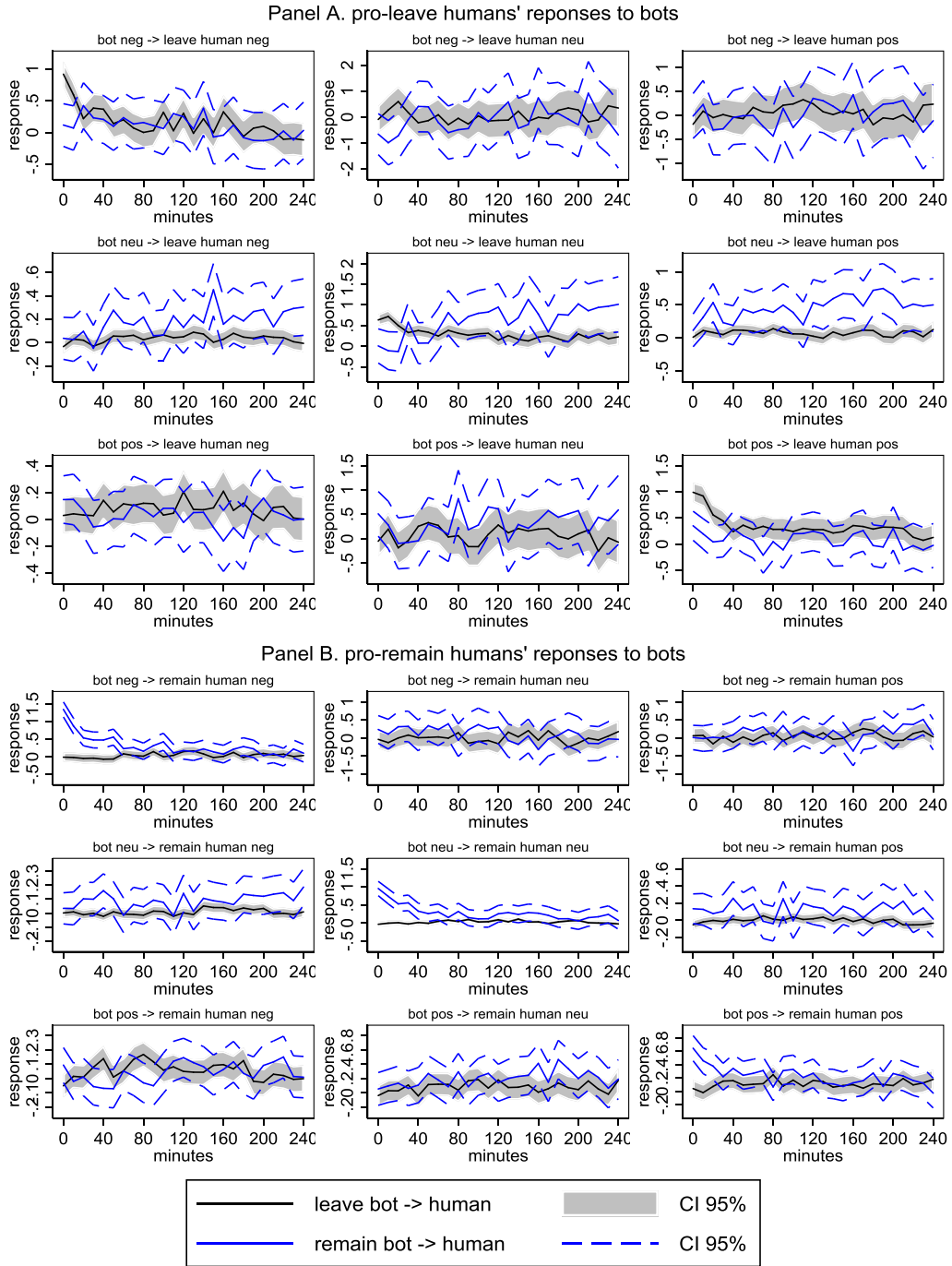
Notes: This figure reports estimate impulse responses of humans to tweets created by bots for the Brexit sample. Panels A and B show the reactions of humans supporting the leave campaign and the remain campaign, respectively. Time (horizontal axis) is in 10-minute intervals. The thin black line refers to the reactions to pro-leave bots while the thick blue line refers to the reactions to pro-remain bots. The grey shaded area and the dashed blue lines indicate 1.96 standard deviation confidence intervals for responses to pro-leave bots and pro-remain bots, respectively.

Figure 9. U.S. (baseline)



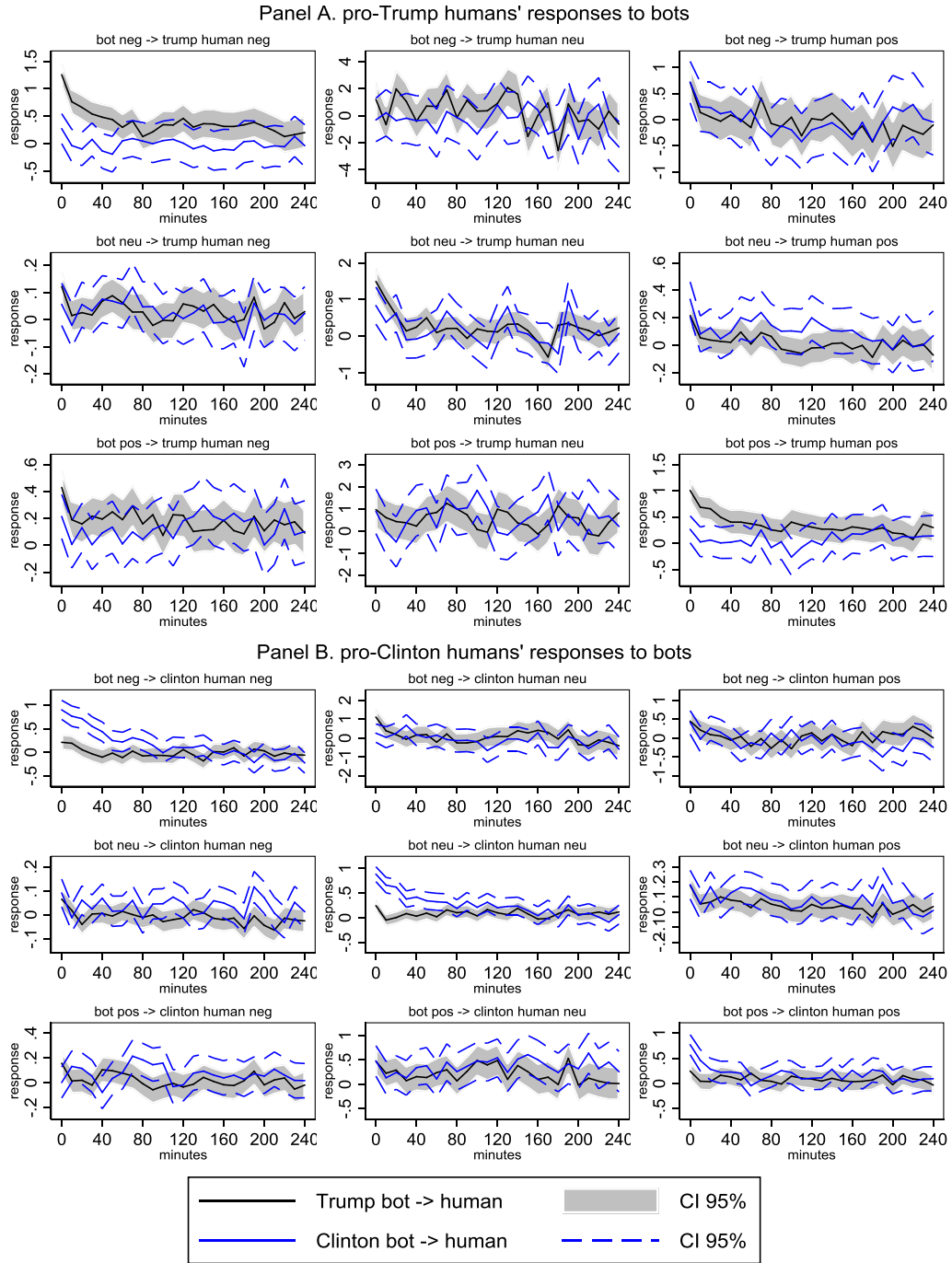
Notes: This figure reports estimate impulse responses of humans to tweets created by bots for the U.S. Election sample. Panels A and B show the reactions of humans supporting Trump and Clinton, respectively. Time (horizontal axis) is in 10-minute intervals. The thin black line refers to the reactions to pro-Trump bots while the thick blue line refers to the reactions to pro-Clinton bots. The grey shaded area and the dashed blue lines indicate 1.96 standard deviation confidence intervals for responses to pro-Trump bots and pro-Clinton bots, respectively.

Figure 10. U.K. (by sentiment)



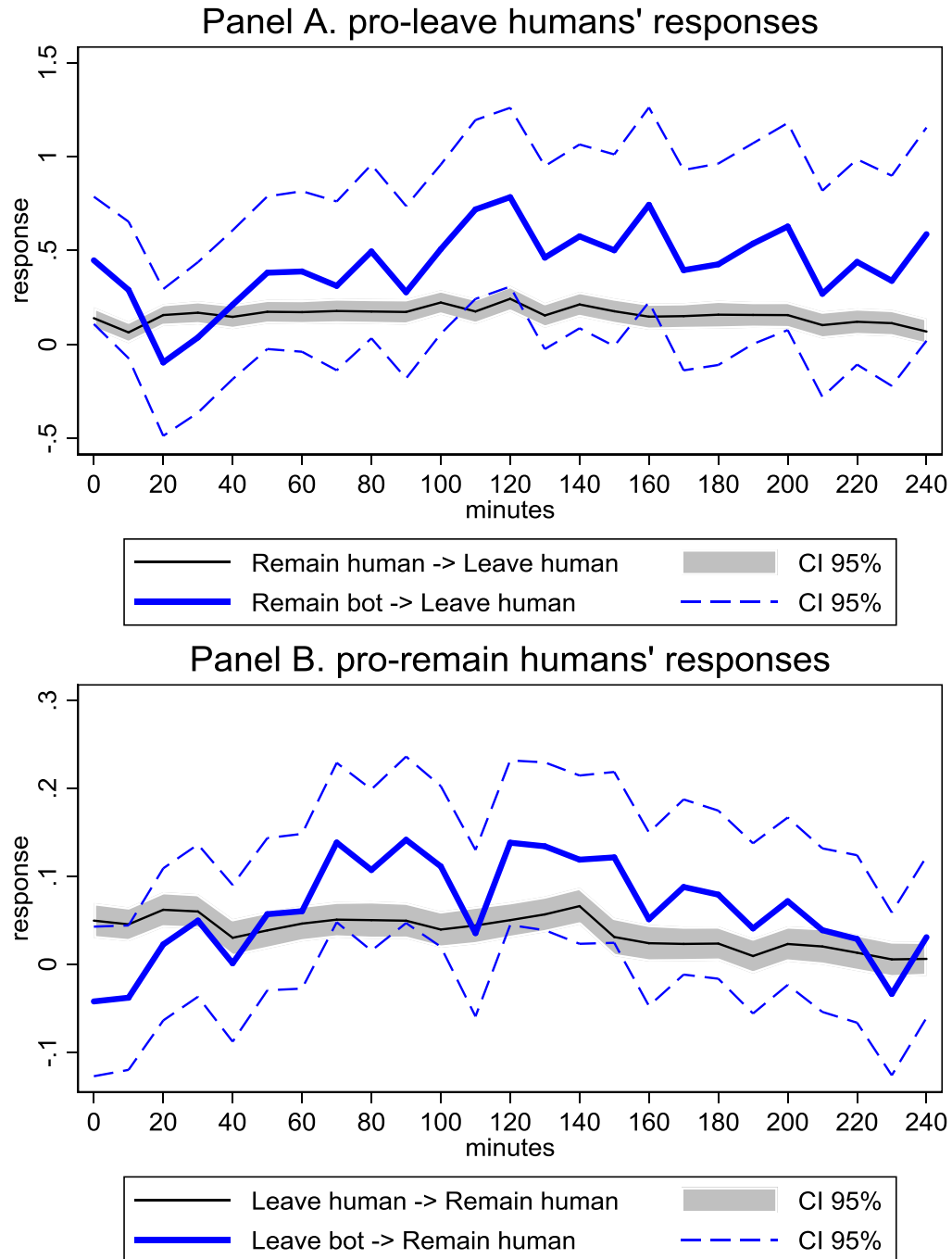
Notes: This figure reports the estimated impulse responses (measured as multipliers) of humans to bots' tweets with different sentiment for the Brexit sample. Panels A and B show reactions of humans supporting the leave campaign and the remain campaign, respectively. Time (horizontal axis) is in 10-minute intervals. The solid black line refers to the reactions to pro-leave bots while the solid blue line refers to the reactions to pro-remain bots. The grey shaded area and the dashed blue lines indicate 1.96 standard deviation confidence intervals for responses to pro-leave bots and pro-remain bots, respectively.

Figure 11. U.S. (by sentiment)



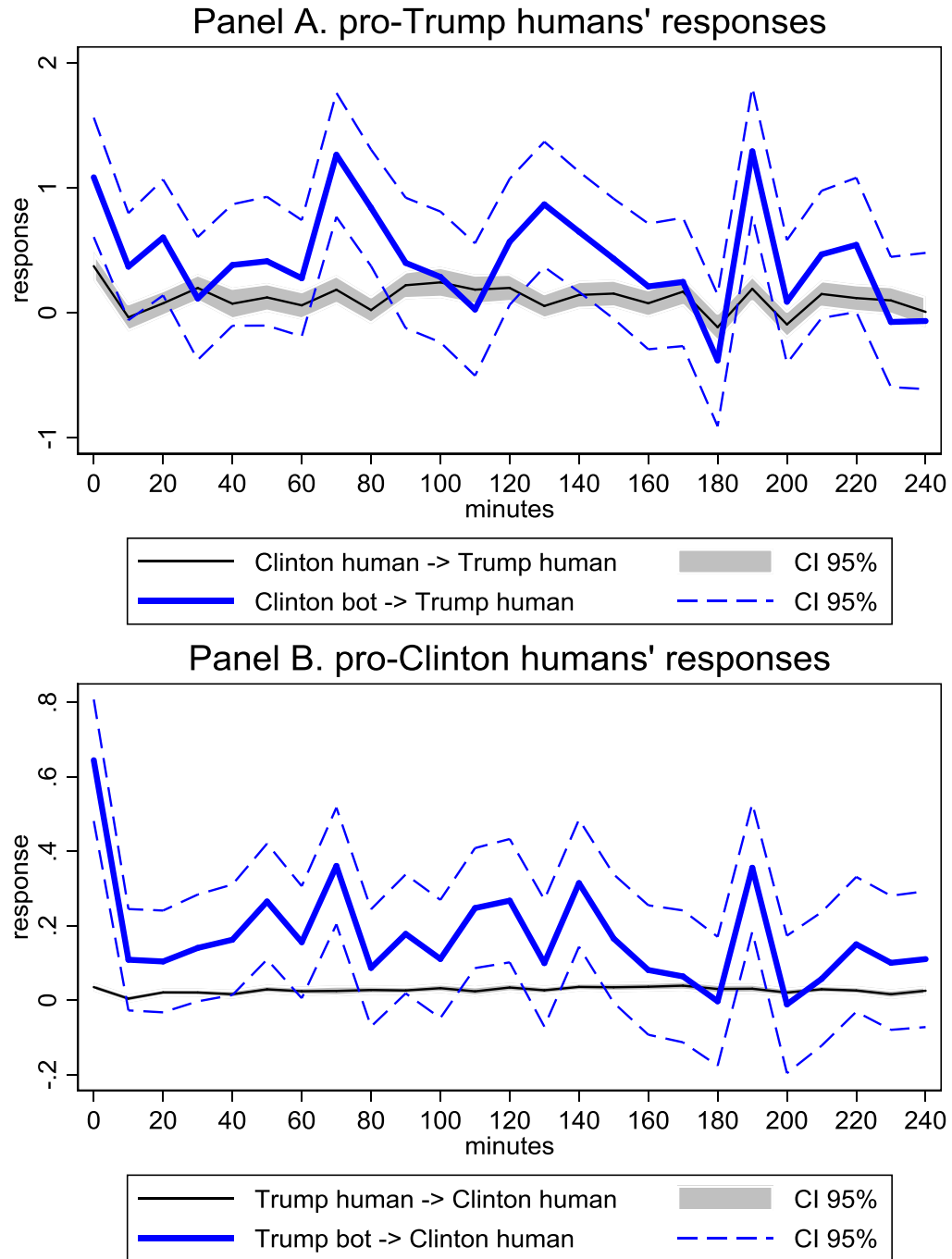
Notes: This figure reports the estimated impulse responses (measured as multipliers) of humans to bots' tweets with different sentiment for the U.S. Election sample. Panels A and B show reactions of humans supporting Trump and Clinton, respectively. Time (horizontal axis) is in 10-minute intervals. The solid black line refers to the reactions to pro-Trump bots while the solid blue line refers to the reactions to pro-Clinton bots. The grey shaded area and the dashed blue lines indicate 1.96 standard deviation confidence intervals for responses to pro-Trump bots and pro-Clinton bots, respectively.

Figure 12. U.K.: humans vs bots in effectiveness to generate tweets of the other side



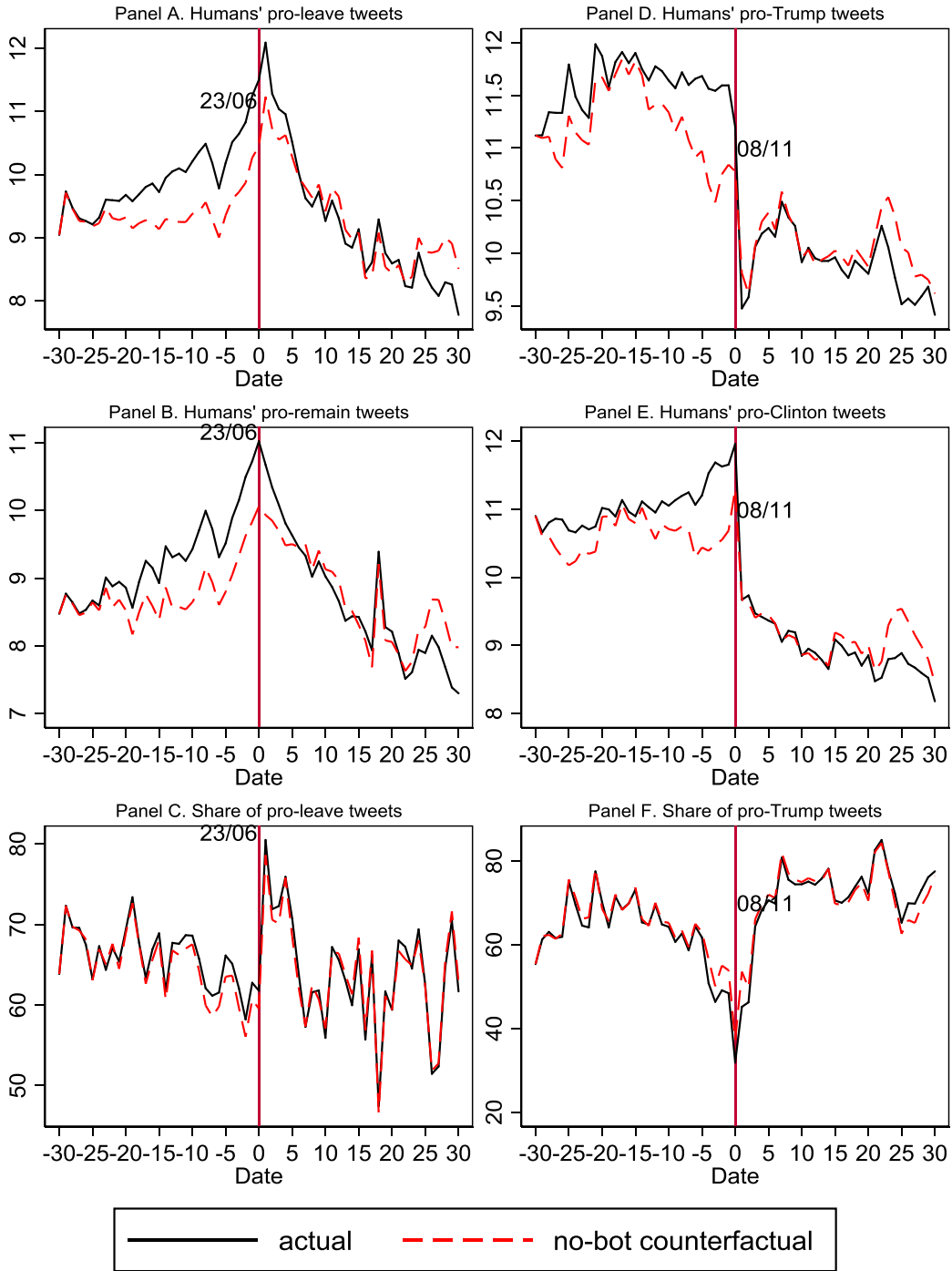
Notes: This figure reports the estimated impulse responses (measured as multipliers) of humans to tweets supporting the opposite campaign for the Brexit sample. Panels A and B show reactions of humans supporting the leave campaign and the remain campaign, respectively. Time (horizontal axis) is in 10-minute intervals. The thin black line refers to the reactions of humans to humans while the thick blue line refers to the reactions of humans to bots. The grey shaded area and the dashed blue lines indicate 1.96 standard deviation confidence intervals for responses to tweets created by humans and bots, respectively.

Figure 13. U.S.: humans vs bots in effectiveness to generate tweets of the other side



Notes: This figure reports the estimated impulse responses (measured as multipliers) of humans to tweets supporting the opposite campaign for the U.S. Election sample. Panels A and B show reactions of humans supporting Trump and Clinton, respectively. Time (horizontal axis) is in 10-minute intervals. The thin black line refers to the reactions of humans to humans while the thick blue line refers to the reactions of humans to bots. The grey shaded area and the dashed blue lines indicate 1.96 standard deviation confidence intervals for responses to tweets created by humans and bots, respectively.

Figure 14. Historical contribution



Notes: This figure reports the historical contribution of bots' traffic to humans' tweet volumes. Panels A, B, and C show bots' contribution to the volumes of pro-leave, pro-remain tweets and the share of pro-leave tweets, respectively. Panels D, E, and F show bots' contribution to the volumes of pro-Trump, pro-Clinton tweets, and the share of pro-Trump tweets, respectively. Time (horizontal axis) is in days. The solid black line refers to the observed human tweet activity while the dashed red line refers to human tweet activity when bots are not present.